# Appendix A <br> Technical Details on the Implementation of the Bi-factor Model 

The estimation equations and expressions for the covariance matrix of the estimates are easily derived using Fisher's identity (Efron 1977; Louis 1982; Glas 1999). The identity plays an important role in the framework of the EM algorithm, which is an algorithm for finding the maximum of a likelihood marginalized over unobserved data. The principle can be summarized as follows. Let $L_{0}(\lambda)$ be the log-likelihood function of parameters $\lambda$ given observed data $x_{0}$, and let $L_{c}(\lambda)$ be the log-likelihood function given both observed data $x_{0}$ and unobserved missing data $x_{m}$. The latter is called the complete data log-likelihood. The interest is in finding expressions for the first-order derivatives of $L_{0}(\lambda)$, say, the expressions for $L_{0}^{\prime}(\lambda)$. Define the first-order derivatives with respect to the complete data log-likelihood as $L_{m}^{\prime}(\lambda)$. Then Fisher's identity entails that $L_{0}^{\prime}(\lambda)$ is equal to the expectation of $L_{m}^{\prime}(\lambda)$ with respect to the posterior distribution of the missing data given the observed data, $p\left(x_{m} \mid x_{0} ; \lambda\right)$, that is,

$$
L_{0}^{\prime}(\lambda)=E\left(L_{m}^{\prime}(\lambda) \mid x_{0}, \lambda\right)=\int L_{m}^{\prime}(\lambda) p\left(x_{m} \mid x_{0}, \lambda\right) d x_{m}
$$

To apply this framework to IRT, a very general definition of an IRT model is adopted. Assume an IRT model is defined by the probability of a response pattern $x_{n}$, which is a function of a, possibly vector-valued, student parameters $\theta_{n}$, and item parameters $a$ and $b$, which are item discrimination and item location parameters of an IRT model. So the IRT model is given by $p\left(x_{n} \mid \theta_{n}, a, b\right)$. Assume further that the student parameter $\theta_{n}$ has a normal density $N\left(\theta_{n} ; \mu_{g(n)}, \Sigma_{g(n)}\right)$ where, again, $g(n)$ is the country to which student $n$ belongs. The key idea is to view the student parameters $\theta_{n}$ as missing data and the item and population parameters $a, b, \mu_{g(n)}$, and $\Sigma_{g(n)}$ as structural parameters $\lambda$ to be estimated. Then the complete data $\log$-likelihood for a student $n$ is

$$
L_{c(n)}(\lambda)=\log p\left(\mathrm{x}_{n} \mid \theta_{n 0}, \theta_{n g}, a, b\right)+\log N\left(\theta_{n 0}, \theta_{n g} ; \mu_{g(n)}, \Sigma_{g(n)}\right)
$$

and so the estimation equations are given by

$$
\begin{gathered}
\frac{\partial L_{0}(\lambda)}{\partial a_{i 0}}=\sum_{n} E\left(\theta_{n 0}\left(\sum_{j=1}^{m_{i}} x_{n i j}-p_{i j}\left(\theta_{n}\right)\right) \mid x_{n}, \lambda\right)=0 \\
\frac{\partial L_{0}(\lambda)}{\partial a_{i g}}=\sum_{n \mid g(n)=g} E\left(\theta_{n g}\left(\sum_{j=1}^{m_{i}} x_{n i j}-p_{i j}\left(\theta_{n}\right)\right) \mid x_{n}, \lambda\right)=0 \\
\frac{\partial L_{0}(\lambda)}{\partial d_{i j}}=\sum_{n}\left[E\left(p_{i j}\left(\theta_{n}\right) \mid x_{n}, \lambda\right)-x_{n i j}\right]=0 \\
\frac{\partial L_{0}(\lambda)}{\partial \mu_{g}}=\sum_{n \mid g(n)=g} \mu_{g}-E\left(\theta_{n 0} \mid x_{n}, \lambda\right)=0 \\
\frac{\partial L_{0}(\lambda)}{\partial \sigma_{g}^{2}}=\sum_{n \mid g(n)=g} \sigma_{g}^{2}-E\left(\theta_{n 0}^{2}-\mu_{g}^{2} \mid x_{n}, \lambda\right)=0
\end{gathered}
$$

where all the expectations are relative to the posterior distribution

$$
p\left(\theta_{n 0}, \theta_{n g} \mid x_{n}, \lambda\right) \propto p\left(\mathrm{x}_{n} \mid \theta_{n 0}, \theta_{n g}, a, b\right) N\left(\theta_{n 0}, \theta_{n g} ; \mu_{g}, \Sigma_{g}\right)
$$

We undertook all calculations using the public domain software package MIRT (Glas 2010). The program uses the EM-algorithm to solve the estimation equations and Gaussian quadrature to evaluate the integrals.

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